# Fastest Expected Time to Mixing for a Markov Chain on a Directed Graph 

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#### Abstract

For an irreducible stochastic matrix $T$, the Kemeny constant $K(T)$ measures the expected time to mixing of the Markov chain corresponding to $T$. Given a strongly connected directed graph $D$, we consider the set $\Sigma_{D}$ of stochastic matrices whose directed graph is subordinate to $D$, and compute the minimum value of $K$, taken over the set $\Sigma_{D}$. The matrices attaining that minimum are also characterised, thus yielding a description of the transition matrices in $\Sigma_{D}$ that minimise the expected time to mixing. We prove that $K(T)$ is bounded from above as $T$ ranges over the irreducible members of $\Sigma_{D}$ if and only if $D$ is an intercyclic directed graph, and in the case that $D$ is intercyclic, we find the maximum value of $K$ on the set $\Sigma_{D}$. Throughout, our results are established using a mix of analytic and combinatorial techniques.


Keywords: Stochastic matrix; Directed graph; Kemeny constant.

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## 1 Introduction and preliminaries

Let $T$ be an irreducible stochastic matrix of order $n$, and denote the stationary distribution of $T$ by $\pi^{T}$. Fix an index $i$ between 1 and $n$, and recall that the Kemeny constant for $T$ is given by $K(T)=\sum_{j=1}^{n} m_{i j} \pi_{j}$, where for each $i, j=1, \ldots, n, m_{i j}$ denotes the mean first passage time from state $i$ to state $j$ (here we take the convention that $m_{i i}=0$ ). It turns out that, remarkably, $K(T)$ is independent of the choice of $i([9])$. Despite its probabilistic formulation, the Kemeny constant can be computed from the eigenvalues of $T$ as follows: denoting the eigenvalues of $T$ by $1 \equiv \lambda_{1}, \lambda_{2}, \ldots, \lambda_{n}$, we have (see [12])

$$
\begin{equation*}
K(T)=\sum_{j=2}^{n} \frac{1}{1-\lambda_{j}} . \tag{1}
\end{equation*}
$$

Indeed, that expression is used (in [8], for example) to show that $K(T) \geq \frac{n-1}{2}$, with equality holding if $T$ happens to be the adjacency matrix of a directed $n$-cycle. Observe that the right hand side of (1) is well-defined for any stochastic matrix $T$ having 1 as a simple (i.e. algebraically and geometrically simple) eigenvalue. Consequently, we slightly extend the definition of the Kemeny constant to the class of stochastic matrices having 1 as a simple eigenvalue, and take $K(T)$ to be given by (1) for all such matrices.

The Kemeny constant admits several interpretations for the Markov chain associated with an irreducible stochastic matrix. In [8], it is shown that $K(T)+1$ coincides with the expected time to mixing for the chain. Here is the idea: let $Y$ be a random variable with probability distribution given by $\pi^{T}$; sample $Y$, say $Y=j$ (with probability $\pi_{j}$ ), and start the chain $\left\{X_{m}\right\}$ at state $X_{0}=i$; define the time to mixing, $M$, to be the minimum $k \geq 1$ such that $X_{k}=j$. It then follows that the expected value for $M$ coincides with $K(T)+1$. In a somewhat different direction, using the fact that $K(T)=\sum_{i=1}^{n} \pi_{i} \sum_{j=1}^{n} m_{i j} \pi_{j}$, the Kemeny constant is interpreted in [12] as the mean first passage time from an unknown initial state to an unknown destination state.

It is also noteworthy that the Kemeny constant provides a measure of the conditioning of the stationary distribution under perturbation of the underlying transition matrix. Specifically, if $T$ and $T+E$ are two irreducible stochastic matrices of order $n$ with stationary distributions $\pi^{T}$ and $\tilde{\pi}^{T}$ respectively, then as shown in [8], we have

$$
\begin{equation*}
\left\|\pi^{T}-\tilde{\pi}^{T}\right\|_{1} \leq K(T)\|E\|_{\infty} . \tag{2}
\end{equation*}
$$

For any stochastic matrix $T$ of order $n$, the directed graph associated with $T, \mathcal{D}(T)$, is the directed graph on vertices labeled $1, \ldots, n$, such that for each $i, j=1, \ldots, n, i \rightarrow j$ is an arc in $\mathcal{D}(T)$ if and only if $t_{i j}>0$. Note that $\mathcal{D}(T)$ carries qualitative information about the Markov chain associated with $T$, since the arcs of $\mathcal{D}(T)$ correspond to the transitions that are possible in a single step of the Markov chain. (We refer the reader to [2] for background on the interplay between square matrices and their directed graphs.) In this paper, we consider the effect of the combinatorial structure of $\mathcal{D}(T)$ on the value of $K(T)$. Specifically, for a strongly connected directed graph $D$ on $n$ vertices, we define the set $\Sigma_{D}$ as follows:

$$
\begin{array}{r}
\Sigma_{D}=\{T \mid T \text { is stochastic and } n \times n \text { and for each } i, j=1, \ldots, n, \\
i \rightarrow j \text { is an arc in } \mathcal{D}(T) \text { only if } i \rightarrow j \text { is an arc in } D\} .
\end{array}
$$

Observe that $\Sigma_{D}$ is a compact, convex set of matrices, whose irreducible members are dense in $\Sigma_{D}$. One of our main results, Theorem 2.6, provides a formula for $\min \left\{K(T) \mid T \in \Sigma_{D}\right\}$, while Theorem 2.13 characterises the matrices yielding that minimum value, thus identifying those transition matrices in $\Sigma_{D}$ that minimise the expected time to mixing. The following example illustrates the scenario that we address in this paper.


Figure 1: Directed graph $D$ for Example 1.1

Example 1.1 Consider the directed graph $D$ shown in Figure 1. A typical irreducible matrix $T \in \Sigma_{D}$ has the form $T=\left[\begin{array}{cccc}0 & 1 & 0 & 0 \\ x & 0 & 1-x & 0 \\ 0 & 1-y & 0 & y \\ 0 & 0 & 1 & 0\end{array}\right]$, where $x, y \in(0,1)$.

It is straightforward to determine that the eigenvalues of such a $T$ are given by $1,-1, \sqrt{x y}$, and $-\sqrt{x y}$. Consequently we find that $K(T)=\frac{1}{2}+\frac{2}{1-x y}$. In particular we have $K(T)>\frac{5}{2}$ for any irreducible $T \in \Sigma_{D}$; note also that $K(T)$ is unbounded from above as $T$ ranges over the irreducible members of $\Sigma_{D}$.

One of the key techniques employed in this paper involves the group inverse of $I-T$, which we now briefly outline. For a stochastic matrix $T$ having 1 as a simple eigenvalue, the singular matrix $I-T$ is known to have a group inverse, $(I-T)^{\#}$, that can be characterised as the unique matrix such that $(I-T)(I-T)^{\#}=(I-T)^{\#}(I-T),(I-T)(I-T)^{\#}(I-T)=(I-T)$, and $(I-T)^{\#}(I-T)(I-T)^{\#}=(I-T)^{\#}$. If the eigenvalues of $T$ are given by $1, \lambda_{2}, \ldots, \lambda_{n}$, then the eigenvalues of $(I-T)^{\#}$ are given by $0, \frac{1}{1-\lambda_{2}}, \ldots, \frac{1}{1-\lambda_{n}}$. In particular, we find that $K(T)=\operatorname{trace}\left((I-T)^{\#}\right)$. If $T$ is a stochastic matrix with 1 as a simple eigenvalue, it follows from Lemma 3.3 of [11] that there is a neighbourhood of $T$ such that $(I-\tilde{T})^{\#}$ is a well-defined continuous function for any stochastic matrix $\tilde{T}$ in that neighbourhood. In particular we see that $K$ is continuous in some neighbourhood of $T$. We refer the interested reader to [3] for further information on generalised inverses.

Throughout the paper, we make use of standard facts on stochastic matrices. The reader may refer to [14] for the necessary background. Background material on directed graphs may be found in [5].

We close this section with a remark on the title of this paper. There is an existing body of work on the so-called fastest mixing Markov chain on a graph (see [1]). Results in that area focus on reversible Markov chains having a specified undirected graph, and on the subdominant eigenvalue of the corresponding transition matrix - i.e. the eigenvalue of second largest modulus after 1 . The object is then to identify the reversible Markov chain respecting that graph which minimizes the modulus of the subdominant eigenvalue of the corresponding transition matrix. Our results in Section 2 bear a philosophical resemblance to those on fastest mixing Markov chains, for we consider an underlying combinatorial structure (a directed graph), and a measure of how quickly a Markov chain mixes $(K(T)+1$ in our case); we then identify those transition matrices that simultaneously respect the combinatorial structure and minimise our measure of mixing time.

## 2 The minimum Kemeny constant on a directed graph

Throughout this section, we take $D$ to be a strongly connected directed graph on $n$ vertices, and we let $k$ denote the length of a longest cycle in $D$. Let $\mu(D)=\inf \left\{K(T) \mid T \in \Sigma_{D}\right.$ and $T$ has 1 as a simple eigenvalue $\}$. We begin with a useful result that leads to an upper bound on $\mu(D)$.

Lemma 2.1 Suppose that $T \in \Sigma_{D}$ has the form

$$
T=\left[\begin{array}{c|c}
C & 0  \tag{3}\\
\hline X & N
\end{array}\right]
$$

where $N$ is nilpotent, and $C$ is the adjacency matrix of a directed cycle of length $\ell$. Then $K(T)=\frac{2 n-\ell-1}{2}$.

Proof. It is straightforward to determine that

$$
(I-T)^{\#}=\left[\begin{array}{c|c}
(I-C)^{\#} & 0 \\
\hline(I-N)^{-1} X(I-C)^{\#}-\frac{1}{\ell}(I-N)^{-1} \mathbf{1 1}^{T} & (I-N)^{-1}
\end{array}\right]
$$

where we use 1 to denote an all ones vector of the appropriate order. Hence $K(T)=K(C)+n-\ell$. From Theorem 3 of [10], we find that each diagonal entry of $(I-C)^{\#}$ is equal to $\frac{\ell-1}{2 \ell}$, so that $K(C)=\frac{\ell-1}{2}$; the conclusion now follows.

Corollary 2.2 We have $\mu(D) \leq \frac{2 n-k-1}{2}$.
Proof. Note that there is a spanning subgraph of $D$ such that each vertex has outdegree 1 , and which contains exactly one directed cycle, of length $k$. Indeed, such a subgraph $\tilde{D}$ can be constructed as follows. Begin by identifying a $k$-cycle in $D$, let $V_{0}$ denote the subset consisting of the vertices on that cycle, and let $A_{0}$ denote the collection of arcs on that cycle. Then, for each $l \geq 0$ such that $\left|\cup_{p=0}^{l} V_{p}\right|<n$, let $V_{l+1}$ denote the set of all vertices in $D$ from which there is an out-arc to some vertex in $V_{l}$; for each $i \in V_{l+1}$, select a single vertex $j_{i} \in V_{l}$ such that $i \rightarrow j_{i}$ is an arc in $D$, and let $A_{l+1}$ denote a collection of arcs $i \rightarrow j_{i}, i \in V_{l+1}$. For some smallest index $m$ we have $\left|\cup_{p=0}^{m} V_{p}\right|=n$, and now we let $\tilde{D}$ be the (spanning) subgraph of $D$ whose arc set is $\cup_{p=0}^{m} A_{p}$.

Let $A$ be the adjacency matrix of such a subgraph $\tilde{D}$, and note then that $A \in \Sigma_{D}$. Observe that $A$ can be written in the form (3), with $C$ as the adjacency
matrix of the directed $k$-cycle. From Lemma 2.1, we have $K(A)=\frac{2 n-k-1}{2}$, from which the conclusion follows.

Our next result shows that if $T$ is a stochastic matrix such that $K(T)$ is not too large, then the non-Perron eigenvalues of $T$ are bounded away from 1 .

Lemma 2.3 Suppose that $A$ is a stochastic matrix of order $n$ having 1 as a simple eigenvalue, and let $\lambda \neq 1$ be an eigenvalue of $A$. If $K(A) \leq n$, then $|1-\lambda| \geq \frac{1-\cos \left(\frac{2 \pi}{n}\right)}{n}$.

Proof. Suppose first that $\lambda \in \mathbb{R}$. In that case, we have $\frac{1}{|1-\lambda|}=\frac{1}{1-\lambda} \leq K(A) \leq n$, so that $|1-\lambda| \geq \frac{1}{n}$, and the desired inequality follows.

Next we suppose that $\lambda$ is complex, say with $\lambda=x+i y$. We then have $n \geq K(A) \geq \frac{1}{1-\lambda}+\frac{1}{1-\bar{\lambda}}=\frac{2(1-x)}{(1-x)^{2}+y^{2}}$. From Theorem 2 of $[6]$ we have $|y| \leq(1-$ $x) \frac{\sin \left(\frac{2 \pi}{n}\right)}{1-\cos \left(\frac{2 \pi}{n}\right)}$, so that $y^{2} \leq(1-x)^{2} \frac{\sin ^{2}\left(\frac{2 \pi}{n}\right)}{\left(1-\cos \left(\frac{2 \pi}{n}\right)\right)^{2}}$. It now follows that $\frac{2(1-x)}{(1-x)^{2}+y^{2}} \geq$ $\frac{1-\cos \left(\frac{2 \pi}{n}\right)}{1-x}$. Hence we find that $|1-\lambda| \geq 1-x \geq \frac{1-\cos \left(\frac{2 \pi}{n}\right)}{n}$.

Next, we show that while $\mu(D)$ is defined as an infimum, it is in fact a minimum.

Lemma 2.4 There is a matrix $S \in \Sigma_{D}$ such that $S$ has 1 as a simple eigenvalue, and $K(S)=\mu(D)$.

Proof. From the definition of $\mu(D)$, we find that there is a sequence of matrices $T_{m} \in \Sigma_{D}$, each with 1 as a simple eigenvalue, such that $K\left(T_{m}\right) \rightarrow \mu(D)$ as $m \rightarrow \infty$. As $\Sigma_{D}$ is compact, there is a subsequence $T_{m_{j}}$ of $T_{m}$ such that $T_{m_{j}}$ converges in $\Sigma_{D}$ as $j \rightarrow \infty$. Denote $\lim _{j \rightarrow \infty} T_{m_{j}}$ by $S$.

Since $K\left(T_{m_{j}}\right) \leq n$ for all sufficiently large $j$, we find from Lemma 2.3 that for all such $j$, and any eigenvalue $\lambda \neq 1$ of $T_{m_{j}},|1-\lambda| \geq \frac{1-\cos \left(\frac{2 \pi}{n}\right)}{n}$. It now follows that the matrix $S$ has 1 as a simple eigenvalue. Thus, the function $K$ is continuous in a neighbourhood of $S$, and so we find that $K(S)=\lim _{j \rightarrow \infty} K\left(T_{m_{j}}\right)=$ $\mu(D)$.

Our next technical result shows that there is a matrix with special structure that minimises $K$.

Lemma 2.5 There is a $(0,1)$ matrix $A \in \Sigma_{D}$ such that 1 is a simple eigenvalue of $A$ and $K(A)=\mu(D)$.

Proof. Appealing to Lemma 2.4, let $S$ be a matrix in $\Sigma_{D}$ having 1 as a simple eigenvalue, such that $K(S)=\mu(D)$. If $S$ is a $(0,1)$ matrix, there is nothing to show, so suppose that some row of $S$ contains at least two positive entries. For concreteness, we take $s_{i p}, s_{i q}>0$, for indices $i, p, q \in\{1, \ldots, n\}$ with $p \neq q$. We claim that there is another matrix in $\Sigma_{D}, \hat{S}$ say, such that $K(\hat{S})=\mu(D)$, and in addition such that $\hat{S}$ has fewer nonzero entries than $S$ does. The conclusion will then follow via an iterative argument.

Let $Q=I-S$, and for each $t \in\left[-s_{i p}, s_{i q}\right]$, let $E_{t}=t e_{i}\left(e_{p}-e_{q}\right)^{T}$. Observe that $S+E_{t} \in \Sigma_{D}$ for all such $t$. Let $\pi^{T}$ denote the stationary distribution for $S$. From Lemma 3.3 of [11], we find that for each $t \in\left[-s_{i p}, s_{i q}\right]$ such that $S+E_{t}$ has 1 as a simple eigenvalue, we have

$$
\left(Q-E_{t}\right)^{\#}=Q^{\#}\left(I-E_{t} Q^{\#}\right)^{-1}-\mathbf{1} \pi^{T}\left(I-E_{t} Q^{\#}\right)^{-1} Q^{\#}\left(I-E_{t} Q^{\#}\right)^{-1}
$$

provided that $I-E_{t} Q^{\#}$ is invertible.
From the Sherman-Morrison formula (see [7] for example) we find that for any $t$ such that $1-t\left(e_{p}-e_{q}\right)^{T} Q^{\#} e_{i} \neq 0$, we have $\left(I-E_{t} Q^{\#}\right)^{-1}=I+$ $\frac{t}{1-t\left(e_{p}-e_{q}\right)^{T} Q^{\#} e_{i}} e_{i}\left(e_{p}-e_{q}\right)^{T} Q^{\#}$. Observe in particular that $\left(I-E_{t} Q^{\#}\right)^{-1} \mathbf{1}=\mathbf{1}$ for any such $t$.

Next, we consider $K\left(S+E_{t}\right)$, and note that for all $t$ such that $|t|$ is sufficiently small, we have

$$
\begin{array}{r}
K\left(S+E_{t}\right)=\operatorname{trace}\left(\left(Q-E_{t}\right)^{\#}\right)= \\
\operatorname{trace}\left(Q^{\#}\right)+\frac{t}{1-t\left(e_{p}-e_{q}\right)^{T} Q^{\#} e_{i}} \operatorname{trace}\left(Q^{\#} e_{i}\left(e_{p}-e_{q}\right)^{T} Q^{\#}\right) \\
-\operatorname{trace}\left(\mathbf{1} \pi^{T}\left(I-E_{t} Q^{\#}\right)^{-1} Q^{\#}\left(I-E_{t} Q^{\#}\right)^{-1}\right) .
\end{array}
$$

Recalling that for any square rank one matrix $a b^{T}$, we have $\operatorname{trace}\left(a b^{T}\right)=b^{T} a$, we find that $\operatorname{trace}\left(Q^{\#} e_{i}\left(e_{p}-e_{q}\right)^{T} Q^{\#}\right)=\left(e_{p}-e_{q}\right)^{T} Q^{\#} Q^{\#} e_{i}$. Also, $\operatorname{trace}\left(\mathbf{1} \pi^{T}(I-\right.$ $\left.\left.\left.E_{t} Q^{\#}\right)^{-1} Q^{\#}\left(I-E_{t} Q^{\#}\right)^{-1}\right)=\pi^{T}\left(I-E_{t} Q^{\#}\right)^{-1} Q^{\#}\left(I-E_{t} Q^{\#}\right)^{-1}\right) \mathbf{1}=\pi^{T}(I-$ $\left.E_{t} Q^{\#}\right)^{-1} Q^{\#} \mathbf{1}=0$. Consequently, we have

$$
\begin{array}{r}
K\left(S+E_{t}\right)=\operatorname{trace}\left(Q^{\#}\right)+\frac{t}{1-t\left(e_{p}-e_{q}\right)^{T} Q^{\#} e_{i}} Q^{\#} e_{i} Q^{\#}\left(e_{p}-e_{q}\right)^{T}= \\
K(S)+\frac{t}{1-t\left(e_{p}-e_{q}\right)^{T} Q^{\#} e_{i}}\left(e_{p}-e_{q}\right)^{T} Q^{\#} Q^{\#} e_{i} .
\end{array}
$$

From the fact that $S$ minimises $K$ over $\Sigma_{D}$, we deduce that $\left(e_{p}-e_{q}\right)^{T} Q^{\#} Q^{\#} e_{i}$ must be zero, otherwise we could select a small (positive or negative) $t$ so
that $K\left(S+E_{t}\right)<\mu(D)$, a contradiction. Consequently we find that for all $t \in\left[-s_{i p}, s_{i q}\right]$ such that $1-t\left(e_{p}-e_{q}\right)^{T} Q^{\#} e_{i} \neq 0, K\left(S+E_{t}\right)=\operatorname{trace}\left(\left(Q-E_{t}\right)^{\#}\right)=$ $\operatorname{trace}\left(Q^{\#}\right)=\mu(D)$. Now select $t_{0}=\left\{\begin{array}{l}s_{i q}, \text { if }\left(e_{p}-e_{q}\right)^{T} Q^{\#} e_{i} \leq 0 \\ -s_{i p}, \text { if }\left(e_{p}-e_{q}\right)^{T} Q^{\#} e_{i}>0\end{array}\right.$, so that $1-t_{0}\left(e_{p}-e_{q}\right)^{T} Q^{\#} e_{i} \geq 1$. Then $S+E_{t_{0}} \in \Sigma_{D}$, has 1 as a simple eigenvalue, has one more zero entry than $S$ does, and satisfies $K\left(S+E_{t_{0}}\right)=\mu(D)$. The conclusion now follows.

Next, we present one of the main results of this paper.

Theorem 2.6 Let $D$ be a strongly connected directed graph on $n$ vertices; denote the length of a longest cycle in $D$ by $k$. Then

$$
\begin{equation*}
\mu(D)=\frac{2 n-k-1}{2} \tag{4}
\end{equation*}
$$

Proof. By Lemma 2.5, there is a $(0,1)$ matrix $A \in \Sigma_{D}$ having 1 as a simple eigenvalue, and such that $K(A)=\mu(D)$. Since $A$ is $(0,1)$ with 1 as a simple eigenvalue, it follows that $A$ can be written in the form (3), where $C$ is the adjacency matrix of a directed cycle, say of length $\ell$, and where $N$ is nilpotent. By Lemma 2.1, we have $\mu(D)=K(A)=\frac{2 n-\ell-1}{2} \geq \frac{2 n-k-1}{2}$. Applying Corollary 2.2, we also have $\mu(D) \leq \frac{2 n-k-1}{2}$, whence $\ell=k$; formula (4) now follows.

Corollary 2.7 Let $T \in \Sigma_{D}$ be irreducible with stationary distribution $\pi^{T}$, and denote the corresponding mean first passage times by $m_{i j}, i, j=1, \ldots, n$. For each index $i=1, \ldots, n$, there is an index $j \neq i$ such that $m_{i j} \geq \frac{2 n-k-1}{2\left(1-\pi_{i}\right)}$.

Proof. From Theorem 2.6, we have $K(T) \geq \frac{2 n-k-1}{2}$. Since $K(T)=\sum_{l \neq i} m_{i l} \pi_{l}$, it follows that $\frac{K(T)}{1-\pi_{i}}$ is a weighted average of the quantities $m_{i l}, l=1, \ldots, n, l \neq i$. The conclusion now follows.

Remark 2.8 Observe that if $T$ is the adjacency matrix of a directed $n$-cycle, then Corollary 2.7 asserts that for each $i=1, \ldots, n$ there is a $j \neq i$ such that $m_{i j} \geq \frac{n}{2}$. If $n$ happens to be even, then we can always find a $j \neq i$ so that in fact $m_{i j}=\frac{n}{2}$.

Example 2.9 Suppose that $T$ is an irreducible tridiagonal stochastic matrix. From the structure of $T$, we find that length of a longest cycle in $\mathcal{D}(T)$ is 2 .

Hence, $K(T) \geq \frac{2 n-3}{2}$ by Theorem 2.6. Thus we have a generalisation of the observations made in Example 1.1.

Our next sequence of results is aimed at characterising the matrices $T \in \Sigma_{D}$ such that $K(T)=\mu(D)$. We begin with a continuity result for minimisers of $K$.

Lemma 2.10 Let $T_{j}$ be a sequence of matrices in $\Sigma_{D}$ such that $K\left(T_{j}\right)=\mu(D)$ for all $j \in \mathbb{N}$. If the sequence $T_{j}$ converges to $S$, then $K(S)=\mu(D)$.
in Proof. Since $K\left(T_{j}\right)=\mu(D)$ for each $j$, we find from Theorem 2.6 and Lemma 2.3 that for any index $j$ and eigenvalue $\lambda \neq 1$ of $T_{j}$, we have $|1-\lambda| \geq \frac{1-\cos \left(\frac{2 \pi}{n}\right)}{n}$. Since the non-Perron eigenvalues of $T_{j}$ are bounded away from 1 , uniformly in $j$, it follows that 1 is a simple eigenvalue of $S$. Hence $K$ is continuous in a neighbourhood of $S$, from which we conclude that $K(S)=\mu(D)$.

Corollary 2.11 Suppose that $T \in \Sigma_{D}$ and that $K(T)=\mu(D)$. Suppose also that there are indices $i, p, q$ with $p \neq q$ such that $t_{i p}, t_{i q}>0$. Letting $S=$ $T+\left(-t_{i p}\right) e_{i}\left(e_{p}-e_{q}\right)^{T}$, we have that $K(S)=\mu(D)$.

Proof. Here we adopt the approach of Lemma 2.5. Let $Q=I-T$, and for each $s \in\left[-t_{i p}, t_{i q}\right]$, let $A_{s}=T+s e_{i}\left(e_{p}-e_{q}\right)^{T}$. Then for each such $s$, we have $K\left(A_{s}\right)=\mu(D)+\frac{s\left(e_{p}-e_{q}\right)^{T} Q^{\#} Q^{\#} e_{i}}{1-s\left(e_{p}-e_{q}\right)^{T} Q^{\#} e_{i}}$, provided that $1-s\left(e_{p}-e_{q}\right)^{T} Q^{\#} e_{i} \neq 0$. As in Lemma 2.5, we deduce that $\left(e_{p}-e_{q}\right)^{T} Q^{\#} Q^{\#} e_{i}=0$, so that $K\left(A_{s}\right)=\mu(D)$ for each $s \in\left[-t_{i p}, t_{i q}\right]$ such that $1-s\left(e_{p}-e_{q}\right)^{T} Q^{\#} e_{i} \neq 0$.

Next, select a sequence $s_{m} \in\left[-t_{i p}, t_{i q}\right]$ such that $1-s_{m}\left(e_{p}-e_{q}\right)^{T} Q^{\#} e_{i} \neq 0$ for all $m \in \mathbb{N}$, and such that $s_{m} \rightarrow-t_{i p}$ as $m \rightarrow \infty$. Then $A_{s_{m}} \rightarrow S$ as $m \rightarrow \infty$, and $K\left(A_{s_{m}}\right)=\mu(D)$ for all $m \in \mathbb{N}$. The conclusion now follows from Lemma 2.10.

The following proposition establishes the combinatorial structure of matrices that minimise $K$ over $\Sigma_{D}$.

Proposition 2.12 Suppose that $A \in \Sigma_{D}$ and that $K(A)=\mu(D)$. Then every cycle in $\mathcal{D}(A)$ has length $k$, and any pair of cycles in $\mathcal{D}(A)$ must intersect.

Proof. We proceed by induction on the number of arcs in $\mathcal{D}(A)$, and note that if $\mathcal{D}(A)$ has just two arcs, then the result is immediate.

Suppose that the conclusion holds for directed graphs with $m \geq 2$ arcs, and that $\mathcal{D}(A)$ has $m+1$ arcs. If each vertex of $\mathcal{D}(A)$ has outdegree one, then since 1 is necessarily a simple eigenvalue of $A$, it follows that $\mathcal{D}(A)$ has just one cycle. Since $K(A)=\mu(D)$, it follows that this cycle must have length $k$, as desired.

Suppose that some vertex of $\mathcal{D}(A)$ has outdegree at least two. Without loss of generality, we assume that $1 \rightarrow i$ and $1 \rightarrow j$ in $\mathcal{D}(A)$. From Corollary 2.11, it follows that we can find $A_{1}, A_{2} \in \Sigma_{D}$ such that $K\left(A_{1}\right)=K\left(A_{2}\right)=\mu(D)$, and such that $\mathcal{D}\left(A_{1}\right)=\mathcal{D}(A) \backslash\{1 \rightarrow i\}$ and $\mathcal{D}\left(A_{2}\right)=\mathcal{D}(A) \backslash\{1 \rightarrow j\}$. Note that every cycle in $\mathcal{D}(A)$ not using the arc $1 \rightarrow i$ is a cycle in $\mathcal{D}\left(A_{1}\right)$, and so by the induction hypothesis, every such cycle has length $k$. On the other hand, any cycle in $\mathcal{D}(A)$ that uses the arc $1 \rightarrow i$ cannot use the arc $1 \rightarrow j$, and so is a cycle in $\mathcal{D}\left(A_{2}\right)$; again by the induction hypothesis, such a cycle must have length $k$. Hence, every cycle in $\mathcal{D}(A)$ has length $k$.

Now select two cycles $C_{1}$ and $C_{2}$ in $\mathcal{D}(A)$. If neither includes the arc $1 \rightarrow i$, then both are in $\mathcal{D}\left(A_{1}\right)$ and hence they must intersect by the induction hypothesis. Evidently if both $C_{1}$ and $C_{2}$ include the arc $1 \rightarrow i$ then they intersect. So, without loss of generality we may assume that $C_{1}$ includes the arc $1 \rightarrow i$, while $C_{2}$ does not. If $C_{2}$ includes the arc $1 \rightarrow j$, then it certainly intersects $C_{1}$, while if $C_{2}$ does not include the arc $1 \rightarrow j$, then both $C_{1}$ and $C_{2}$ are in $\mathcal{D}\left(A_{2}\right)$. Again the induction hypothesis applies, and we find that $C_{1}$ and $C_{2}$ intersect. That completes the induction step, and the conclusion follows.

Recall that an irreducible stochastic matrix $T$ is periodic with period $m$ if the greatest common divisor of the cycle lengths in $\mathcal{D}(T)$ is equal to $m$. In that case, the vertices of $\mathcal{D}(T)$ can be partitioned into subsets $S_{1}, \ldots, S_{m}$ such that $i \rightarrow j$ is an arc in $\mathcal{D}(T)$ only if there is an index $\ell=1, \ldots, m$ such that $i \in S_{\ell}$ and $j \in S_{\ell+1}$ (with the convention that $S_{m+1} \equiv S_{1}$ ). These subsets $S_{1}, \ldots, S_{m}$ are known as the cyclically transferring classes for $T$.

We are now in a position to characterise the matrices that minimise $K$ over $\Sigma_{D}$.

Theorem 2.13 Suppose that $A \in \Sigma_{D}$. We have $K(A)=\mu(D)$ if and only if $A$ can be written in the form

$$
A=\left[\begin{array}{c|c}
A_{0} & 0  \tag{5}\\
\hline X & N
\end{array}\right]
$$

where $N$ is nilpotent (or empty in the case that $k=n$ ) and where $A_{0}$ is irreducible, and $k$-cyclic with one of its cyclically transferring classes of cardinality one.

Proof. Suppose that $K(A)=\mu(D)$; then $A$ has 1 as a simple eigenvalue, and it follows that we may write $A$ as

$$
A=\left[\begin{array}{c|c}
A_{0} & 0 \\
\hline X & N
\end{array}\right]
$$

where $A_{0}$ is irreducible and the spectral radius of $N$ is strictly less than 1 . By Proposition 2.12, all cycles of $\mathcal{D}(A)$ have length $k$, and any two cycles intersect. Hence, all cycles of $\mathcal{D}\left(A_{0}\right)$ have length $k$, and any two cycles intersect; applying Theorem 6.2 of [4], we thus find that $A_{0}$ must be $k$-cyclic with one of its cyclically transferring classes having cardinality one.

It is straightforward to see that $K(A)=\operatorname{trace}\left(\left(I-A_{0}\right)^{\#}\right)+\operatorname{trace}\left((I-N)^{-1}\right)$. Suppose for concreteness that $A_{0}$ is $m \times m$; from the structure of $A_{0}$, we find that its eigenvalues are $e^{\frac{2 \pi i j}{k}}, j=0, \ldots, k-1$, and 0 with algebraic multiplicity $m-k$. Hence $\operatorname{trace}\left(\left(I-A_{0}\right)^{\#}\right)=\frac{k-1}{2}+m-k$. Note also that trace $((I-$ $N)^{-1}$ ) $\geq n-m$, with equality holding only if $N$ is nilpotent. Thus we have $\frac{2 n-k-1}{2}=\frac{k-1}{2}+m-k+\operatorname{trace}\left((I-N)^{-1}\right) \geq \frac{k-1}{2}+m-k+n-m=\frac{2 n-k-1}{2}$. We thus conclude that $N$ must be nilpotent, as desired.

The converse is readily established.

Remark 2.14 Suppose that $T$ is an irreducible stochastic matrix of order $n$. From Theorem 2.6, we recover the known result that $K(T) \geq \frac{n-1}{2}$, while from Theorem 2.13, we find that $K(T)=\frac{n-1}{2}$ if and only if $T$ is the adjacency matrix of a directed $n$-cycle.

Remark 2.15 Let $D$ be a strongly connected directed graph. It is interesting to note that any matrix in $\Sigma_{D}$ that minimises the Kemeny constant necessarily has a subdominant eigenvalue of modulus 1 . Thus we see that by using $K(T)+1$ as a measure for the time to mixing, we obtain very different results than by using the modulus of the subdominant eigenvalue as a measure of the time to mixing.

## 3 An upper bound for intercyclic directed graphs

In light of the lower bound on $K$ established in Theorem 2.6, it is natural to wonder about the structure of the directed graphs $D$ such that $K(T)$ is bounded from above as $T$ ranges over $\Sigma_{D}$. In this section, we address that question. We begin with a useful observation.

Remark 3.1 It is shown in Lemma 6.1 of [4] that if $D$ contains a pair of vertex-disjoint cycles, then $K(T)$ is not bounded from above as $T$ ranges over the irreducible matrices in $\Sigma_{D}$.

Recall that a directed graph is intercyclic if it has the property that any pair of its cycles intersect. A complete characterisation of this class of directed graphs is given in [13]. Observe that by Remark 3.1, if a directed graph $D$ has the property that $K(T)$ is bounded from above as $T$ ranges over the irreducible members of $\Sigma_{D}$, then necessarily $D$ must be intercyclic.

The following technical result will be useful in the sequel.
Lemma 3.2 Let $D$ be an intercyclic directed graph. Then for any matrix $T \in$ $\Sigma_{D}$, the number 1 is a simple eigenvalue of $T$.

Proof. Fix $T \in \Sigma_{D}$; for each cycle $C$ in $\mathcal{D}(T)$, let $w(C)$ denote the product of the entries in $T$ corresponding to the arcs of $C$. Since any pair of cycles in $\mathcal{D}(T)$ intersect, it follows that the characteristic polynomial of $T$ can be written as $\operatorname{det}(\lambda I-T)=\lambda^{n}-\sum_{C \in \mathcal{D}(T)} w(C) \lambda^{n-|C|}$, where the sum is taken over all cycles $C \in \mathcal{D}(T)$, and where $|C|$ denotes the number of vertices on the cycle $C$. It now follows from Descartes' rule of signs that $\operatorname{det}(\lambda I-T)$ has precisely one positive root, which is necessarily equal to 1 . Hence 1 is a simple eigenvalue of $T$.

Lemma 3.2 leads to a continuity result for $K$.
Corollary 3.3 Let $D$ be an intercyclic directed graph. Then $K$ is a continuous function on $\Sigma_{D}$; in particular, there is a matrix $A \in \Sigma_{D}$ such that $K(A)=$ $\max \left\{K(T) \mid T \in \Sigma_{D}\right\}$.

Proof. Let $T$ be a stochastic matrix with 1 as a simple eigenvalue. Then $(I-T)^{\#}$ is continuous in a neighbourhood of $T$, and hence so is $K(T)=\operatorname{trace}\left((I-T)^{\#}\right)$. The other conclusion follows readily.

Here is the main result of this section.
Theorem 3.4 Suppose that $D$ is an intercyclic directed graph, and let $g$ denote the length of a shortest cycle in $D$. Then $\max \left\{K(T) \mid T \in \Sigma_{D}\right\}=\frac{2 n-g-1}{2}$.

Proof. By Lemma 3.3, $K$ attains its maximum value on $\Sigma_{D}$. Arguing as in Lemma 2.5 , it is readily shown that in fact there is a $(0,1)$ matrix $A$ in $\Sigma_{D}$ for which $K(A)$ is maximum. As $A$ has 1 as a simple eigenvalue and is $(0,1)$, it follows that $A$ can be written in the form (3), where $C$ is the adjacency matrix of a directed cycle of length $\ell$, say. It then follows that $K(A)=\frac{2 n-\ell-1}{2} \leq \frac{2 n-g-1}{2}$. On the other hand, we can readily produce a matrix $T$ in $\Sigma_{D}$ such that $\mathcal{D}(T)$ contains a single cycle of length $g$, and for which $K(T)=\frac{2 n-g-1}{2}$. Consequently, it must be the case that $\max \left\{K(T) \mid T \in \Sigma_{D}\right\}=\frac{2 n-g-1}{2}$, as desired.


Figure 2: Directed graph $D$ for Example 3.5

Example 3.5 We close the paper with an example that illustrates the results of this section. Consider the directed graph $D$ shown in Figure 2. It is straightforward to see that $D$ is intercyclic (since vertex 7 is on every cycle), and that the shortest and longest cycle lengths are 3 and 4 , respectively.

Suppose that we have a matrix $T \in \Sigma_{D}$. Then there are parameters $x, y, a \in$ $[0,1]$ such that

$$
T=\left[\begin{array}{ccccccc}
0 & 0 & 1-x & 0 & 0 & x & 0 \\
0 & 0 & 0 & 1-y & y & 0 & 0 \\
0 & 0 & 0 & 0 & 1 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 1 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 1 \\
0 & 0 & 0 & 0 & 0 & 0 & 1 \\
a & 1-a & 0 & 0 & 0 & 0 & 0
\end{array}\right]
$$

Let $U$ be the submatrix of $I-T$ consisting of its first six columns, and let $V$ be the $6 \times 7$ matrix $V=[I \mid \mathbf{- 1}]$. We find from Theorem 7.8.2 of [3] that $(I-T)^{\#}=U(V U)^{-2} V$, from which it follows that $K(T)=\operatorname{trace}\left(U(V U)^{-2} V\right)=$ $\operatorname{trace}\left((V U)^{-1}\right)$. A direct computation now shows that $K(T)=3+\frac{6}{4-a x-(1-a) y}$. Consequently, we find that $K(T) \geq \frac{9}{2}=\frac{14-4-1}{2}$, with equality holding if and only if $a x+(1-a) y=0$, while $K(T) \leq 5=\frac{14-3-1}{2}$, with equality holding if and only if $a x+(1-a) y=1$.

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